

A two-parameter hierarchy of probabilities, the action of the shift on the R_k 's, and an approximation of local time.

1. Motivations.

It seems interesting to introduce the two-parameter family:

$$P_{k,m}^\# = \frac{c_{k,m}}{(\sqrt{\tilde{\sigma}_1})^k (\sqrt{V_1(\tilde{\sigma}_1)})^m} \cdot P^\#$$

In my document (7), I have looked uniquely at $P_{k,0}^\#$, and found an "explicit" density for $P_{k,0}^\# | R_p$ with respect to $P_{k,0}^\# | R_p$.

Here, I shall look at the same quantities for $P_{0,m}^\#$, and apply this to obtain a martingale proof of the following result:

$$(1) \quad \lim_{n \rightarrow \infty} \left(\frac{\pi}{2} n^2 V_n(t) \right) = L_t^2, \text{ uniformly on compacts.}$$

Once this is done, we will look again at "explicit" computations for $P_{k,m}^\# | R_p$ with respect to $P_{k,m}^\# | R_p$.

2. Some absolute continuity relationships.

We first remark that, for every $m > 0$, the sequence:

$$(2) \quad \left\{ \Delta_n^{(m)} \stackrel{\text{def}}{=} \frac{(m+n)!}{m! n!} (R_1 \cdots R_m)^{\frac{m}{2}} \equiv \frac{(m+n)!}{m! n!} \left(\frac{V_{m+1}(1)}{V_1(1)} \right)^{m/2}, m \geq 1 \right\}$$

is a $(P, (R_m))$ martingale (Proof: Immediate from the knowledge of the distribution of the R_m 's).

As a Corollary of (5), we deduce the convergence result (1) immediately, ~~but~~ by taking e.g. : $m=2$ (we also use Dirichlet's thm. to obtain the uniform convergence on compacts !!).

Let us look now closer at (5), and write this identity in the following form:

$$(5') \quad E \left[\left(\frac{L_1}{\sqrt{V_1(1)}} \right)^m \mid \mathcal{R}_p \right] = \frac{(m+p)!}{p!} \left(\frac{\pi}{2} \frac{V_{p+1}(1)}{V_1(1)} \right)^{m/2}$$

Thus, we have obtained the following

Thm: For every p , the $(p+1)$ random variables R_1, R_2, \dots, R_p
and $\left(\frac{L_1}{\sqrt{V_{p+1}}} \right)$ are independent, and, moreover:

$\sqrt{\frac{2}{\pi}} \cdot \frac{L_1}{\sqrt{V_{p+1}}} \stackrel{\text{(law)}}{=} Z_{p+1}$ (: gamma variable with parameter $(p+1)$).

Quelques remarques sur le document (10).

1. La propriété : $\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}} \stackrel{\text{(loi)}}{=} Z_{p+1}$ est bien en accord avec les

identités en loi de l'algèbre beta-gamma ; en effet, on peut écrire :

$$\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_p}} = \left(\frac{\sqrt{V_{p+1}}}{\sqrt{V_p}} \right) \left(\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}} \right),$$

et, dans le membre de droite, $\sqrt{R_p} \equiv \frac{\sqrt{V_{p+1}}}{\sqrt{V_p}}$ a pour loi $Z_{p,1}$ et est indépendante de $\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}}$, qui a pour loi Z_{p+1} .

On retrouve donc bien ainsi l'identité : $Z_p \stackrel{\text{(loi)}}{=} Z_{p,1} Z_{p+1}$.

* i.e. le second
lim. qui figure en
(10).

2. Dans le même esprit, on déduit du même théorème la propriété suivante :
les variables $\left(\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}} - \sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_p}}, p \geq 1 \right)$

~~est~~ sont indépendantes, et équidistribuées, avec pour loi commune la loi exponentielle de paramètre 1.

(C'est toujours une conséquence du théorème, et de l'algèbre beta-gamma)

3. Considérons maintenant la relation (8), p. 6, du document (7), que je réécris sous la forme suivante :

$$(1) \quad E \left[\sqrt{2T} \left(\left(\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}} \right) \sqrt{V_{p+1}} \right)^k \mid \mathcal{R}_p \right]$$

$$= E \left[(Z_{1+p})^k \right] \int_0^{\infty} dx \left(\frac{x e^{x^2/2}}{\psi(x)} \right)^k \gamma_{a,p}(x)$$

Because of this equality, it is now tempting to think that, in fact:

$$(2?) \quad \left(\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}} \right) \text{ is independent of } \sigma \{ \mathcal{R}_{p+1}, V_{p+1} \} \equiv \sigma \{ V_1, V_2, \dots, V_{p+1} \}$$

and that the identity (1) may be rewritten as:

$$(3?) \quad E \left[\left(\sqrt{2T} \sqrt{V_{p+1}} \right)^k \mid \mathcal{R}_{p+1} \right] = \int_0^\infty dx \left(\frac{x e^{x^2/2}}{\Psi(x)} \right)^k \gamma_{q,p}(x).$$

However, (2?) is certainly wrong for the 2 following reasons:

(i) if it were true, then we would deduce from the equality

$$L_1 = \left(\frac{L_1}{\sqrt{V_{p+1}}} \right) \sqrt{V_{p+1}}$$

the identity in law: (4?) $|N| \stackrel{(law)}{=} \left(\sqrt{\frac{\pi}{2}} Z_{1+p} \right) \sqrt{V_{p+1}}$

It should be fairly easy to show that this does not hold.

(ii) Other contradictions arise from (3?).

Indeed, the function: $u(x) = \frac{x e^{x^2/2}}{\Psi(x)}$ is

$$(4) \quad \frac{x e^{x^2/2}}{\Psi(x)} = \frac{x}{\tilde{\Psi}(x)} = \frac{1}{\sqrt{\frac{\pi}{2}} + \int_x^\infty \frac{du}{u^2} \exp(-\frac{u^2}{2})}$$

[Pour la correction du 2 voir le doc. (6), p. 3]

and, as x varies in $]0, \infty[$, the function $u(x)$ varies between 0 and $\sqrt{\frac{2}{\pi}}$.
But, on the left hand side of (3?), the variable $\sqrt{2T} \sqrt{V_{p+1}}$ varies on the entire half-line.

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(10⁴)

The "full" computations for the 2-parameter "hierarchy" of probabilities. (10⁴) 1

I combine the computation of

$$(1) \quad E_{k,m}^{\#} (F_{\nu}) \equiv E^{\#} \left(\frac{c_{k,m} F_{\nu}}{(\sqrt{z_1})^k (\sqrt{V_1(z_1)})^m} \right)$$

which were done for $k=0$ in Doc (10'), and for $m=0$ in Doc. (7).

Here, I shall follow closely the computations made in Doc. (7); the notation $(j)_7$ refers to the equation (j) in Doc(7).
We have:

$$E_{k,m}^{\#} (F_{\nu}) = c_{k,m} E^{\#} \left(\frac{1}{(V_1(z_1))^{\frac{k+m}{2}}} \frac{F_{\nu}(R'_1, n \geq 1)}{(1 + R'_1 + R'_1 R'_2 + \dots + \dots)^{k/2}} \right)$$

$$= c_{k,m} \lim_{n \rightarrow \infty} E \left[F_{\nu} \frac{(n \sqrt{\frac{\pi}{2}})^{k+m} (R_1 \dots R_n)^{\frac{k+m}{2}}}{(1 + R_1 + R_1 R_2 + \dots)^{k/2}} \right]$$

[this generalizes (1)₇]

Now, let $l = k+m$, and define: $c'_{k,m} = c_{k,m} \left(\sqrt{\frac{\pi}{2}}\right)^l (l!).$

Reading Doc. (7), bottom of p.3, we get:

$$E_{k,m}^{\#} (F_{\nu}) = c'_{k,m} E \left[F(R_{l+1}, R_{l+2}, \dots, R_{l+p}) \frac{1}{(\sum_{l+1})^{k/2}} \right],$$

and, from the top of p.4, Doc. (7):

(10'') 2)

$$E \left[\frac{1}{\left(\sum_{l+1} \right)^{k/2}} \middle| R_{l+p} \right]$$

$$= \frac{1}{\Gamma\left(\frac{k}{2}\right)} \int_0^{\infty} dt \, t^{\left(\frac{k}{2}-1\right)} e^{-t(1+R_{l+1}+(R_{l+1}R_{l+2})+\dots+(R_{l+1}\dots R_{l+p}))} \varphi_{l+p+1}(tR_{l+1}\dots R_{l+p})$$

$$= \frac{1}{\Gamma\left(\frac{k}{2}\right) (R_{l+1}\dots R_{l+p})^{k/2}} \int_0^{\infty} dt \, t^{\frac{k}{2}-1} e^{-t\theta_p(R_{l+1}, \dots, R_{l+p})} \varphi_{l+p+1}(t)$$

We now replace the vector $(R_{l+1}, \dots, R_{l+p})$ by (R_1, \dots, R_p) with the help of (2)₇. Thus, we obtain:

$$(2) \quad E_{k,m}^{\#}(F_p) = c'_{k,m} E \left[\frac{(l+p)!}{l! p!} \frac{1}{\Gamma\left(\frac{k}{2}\right)} F_p(R_1, \dots, R_p)^{m/2} \int_0^{\infty} dt \, t^{\frac{k}{2}-1} e^{-t\theta_p(R_1, \dots, R_p)} \varphi_{l+p+1}(t) \right]$$

In particular, we have obtained the following generalization of (3')₇:

$$(3) \quad P_{k,m}^{\#} | R_p = d_{k,m} \left(\theta_p(R_1, \dots, R_p); (R_1, \dots, R_p) \right) \cdot P | R_p$$

where:

$$d_{k,m}(a, b) = c'_{k,m} \frac{(l+p)! \, b^{m/2}}{l! \, p! \, \Gamma\left(\frac{k}{2}\right)} \int_0^{\infty} dt \, t^{\frac{k}{2}-1} e^{-ta} \varphi_{l+p+1}(t)$$

$$l = k+m, \text{ and } c'_{k,m} = c_{k,m} \left(\frac{\sqrt{\pi}}{2} \right)^l (l!).$$

Some consequences of formula (3):

By comparing (1) and (3), and using again the admissibility of \mathcal{E}_1 , we learn the following:

(in the sequel, we note: $a = \theta_p(R_1, \dots, R_p)$, and $b = R_1 \dots R_p \equiv \frac{V_{p+1}}{V_1}$)

$$E \left[c_{k,m} \frac{(L_1)^l}{(\sqrt{V_1})^m} \mid \mathcal{R}_\mu \right] = c'_{k,m} \frac{(l+p)! b^{m/2}}{l! p! \Gamma(\frac{k}{2})} \int_0^\infty dt t^{\frac{k}{2}-1} e^{-ta} \varphi_{l+p+1}(t)$$

hence:

$$E \left[\frac{(L_1)^l}{(\sqrt{V_1})^m} \mid \mathcal{R}_\mu \right] = \left(\sqrt{\frac{\pi}{2}}\right)^l \frac{(l+p)! b^{m/2}}{p! \Gamma(\frac{l-m}{2})} \int_0^\infty dt t^{\frac{l-m}{2}-1} e^{-ta} \varphi_{l+p+1}(t)$$

Writing $V_1 = (V_{p+1}) \left(\frac{1}{b}\right)$, we obtain:

$$E \left[\frac{(L_1)^l}{(\sqrt{V_{p+1}})^m} \mid \mathcal{R}_\mu \right] = \left(\sqrt{\frac{\pi}{2}}\right)^l \frac{(l+p)!}{p! \Gamma(\frac{l-m}{2})} \int_0^\infty dt t^{\frac{l-m}{2}-1} e^{-ta} \varphi_{l+p+1}(t)$$

which, using $l = k+m$ again, φ may be written as:

$$(4) E \left[\left(\sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}}\right)^l (\sqrt{V_{p+1}})^k \mid \mathcal{R}_\mu \right] = \frac{(l+p)!}{p! \Gamma(\frac{k}{2})} \int_0^\infty dt t^{\frac{k}{2}-1} e^{-ta} \varphi_{l+p+1}(t)$$

We now transform (4) by making the change of variables $t = \frac{x^2}{2}$, and we use:

$$\varphi_n\left(\frac{x^2}{2}\right) = \frac{e^{\frac{x^2}{2}(n-1)}}{(\Psi(x))^n}$$

Thus, we obtain, by writing: $X_{p+1} \equiv \sqrt{\frac{2}{\pi}} \frac{L_1}{\sqrt{V_{p+1}}}$:

$$(4') E \left[(X_{p+1})^l (\sqrt{V_{p+1}})^k \mid \mathcal{R}_\mu \right] = \frac{(l+p)!}{p! \Gamma(\frac{k}{2})} \int_0^\infty dx \frac{x^{k-1}}{2^{\frac{k}{2}-1}} e^{-\frac{x^2}{2}} \frac{e^{\frac{x^2}{2}(l+p)}}{(\Psi(x))^{l+p+1}}$$

Now, recall the two following facts:

a) $X_{p+1} \stackrel{\text{(law)}}{=} Z_{p+1}$, and X_{p+1} is independent of \mathcal{R}_μ ;

b) $E \left[(Z_{p+1})^l \right] = \frac{(l+p)!}{p!}$.

Using a) and b), we can write the following variants of (4'):

$$(5) \quad E \left[f(X_{p+1}) (\sqrt{V_{p+1}})^k \mid \mathcal{R}_p \right] \\ = \int_0^\infty dx \frac{x^{k-1}}{2^{\frac{k}{2}-1} \Gamma(\frac{k}{2})} \frac{e^{-\frac{x^2}{2}(a-p)}}{(\psi(x))^{p+1}} E \left[f \left(z_{p+1} \frac{e^{x/2}}{\psi(x)} \right) \right]$$

Denote by $g_{p+1}(z)$ the density of z_{p+1} . Now, using a) above, we easily obtain:

$$g_{p+1}(z) E \left[(\sqrt{V_{p+1}})^k \mid \mathcal{R}_p, X_{p+1}=z \right] \\ = \int_0^\infty dx \left(\frac{x^{k-1}}{2^{\frac{k}{2}-1} \Gamma(\frac{k}{2})} \right) \frac{1}{(\psi(x))^{p+1}} \frac{e^{-\frac{x^2}{2}(a-p)}}{\psi(x) e^{-\frac{x^2}{2}}} g_{p+1}(\psi(x) e^{-\frac{x^2}{2}} z) \\ \text{since } g_{p+1}(z) = \frac{1}{p!} z^p e^{-z}, \text{ this simplifies to give:}$$

$$(6) \quad e^{-z} E \left[(\sqrt{V_{p+1}})^k \mid \mathcal{R}_p, X_{p+1}=z \right] = \int_0^\infty dx \left(\frac{x^{k-1}}{2^{\frac{k}{2}-1} \Gamma(\frac{k}{2})} \right) \frac{1}{(\psi(x))^{p+1}} e^{\left(\frac{x^2}{2}(1+a) + z \psi(x) e^{-\frac{x^2}{2}} \right)}$$

We now introduce T an exponential variable, which is independent of Brownian motion, with parameter 1. Thus, we deduce from (6) that:

$$(7) \quad e^{-z} E \left[(\sqrt{2TV_{p+1}})^k \mid \mathcal{R}_p, X_{p+1}=z \right] \\ = k \int_0^\infty dx x^{k-1} e^{-\frac{x^2}{2}(1+a)} e^{-z \psi(x) e^{-\frac{x^2}{2}}}, \text{ so that, using} \\ \text{integration by parts:}$$

$$(8) \quad e^{-z} E \left[(\sqrt{2TV_{p+1}})^k \mid \mathcal{R}_p, X_{p+1}=z \right] = \int_0^\infty dx x^k \left\{ -\frac{d}{dx} \left(e^{-\frac{x^2}{2}(1+a)} e^{-z \psi(x) e^{-\frac{x^2}{2}}} \right) \right\}$$

Therefore, we have obtained:

$$(9) \quad e^{-z} \mathbb{P} \left(\sqrt{2T} V_{p+1} \in dx \mid \mathcal{R}_p, X_{p+1} = z \right) = dx \left\{ -\frac{d}{dx} \left(e^{-\frac{x^2}{2}(1+a) - z\psi(x)} e^{-\frac{x^2}{2}} \right) \right.$$

Now, we have:

$$e^{-z} \mathbb{P} \left(2TV_{p+1} \geq x^2 \mid \mathcal{R}_p, X_{p+1} = z \right) = e^{-\frac{x^2}{2}(1+a) - z\psi(x)} e^{-\frac{x^2}{2}},$$

and, finally:

$$(10) \quad e^{-z} \mathbb{E} \left[\exp \left(-\frac{x^2}{2V_{p+1}} \right) \mid \mathcal{R}_p, X_{p+1} = z \right] = e^{-\frac{x^2}{2}(1+a) - z\psi(x)} e^{-\frac{x^2}{2}}.$$

a formula which, of course, can be found directly from Michael's paper...
but, at the end of this story, there may be some interest in seeing
how the computations for the 2-parameter hierarchy
are related to the structure of the sequence:
($V_1, V_2, \dots, V_n, \dots$),

which, in the end, is described by the following

Theorem:

$$\text{Define } R_n \equiv \frac{V_{n+1}}{V_n} \quad (n \geq 1).$$

Then, the law of the sequence ($V_n, n \geq 1$) is characterized by the 3 following
properties: (i) the random variables (R_n) are independent, and for each n , $R_n \stackrel{\text{law}}{=} U^{2/n}$.

Now, fix $p \geq 1$. Then:

(ii) $X_{p+1} \stackrel{\text{def}}{=} \sqrt{\frac{\pi}{2}} \frac{L_1}{\sqrt{V_{p+1}}}$ is independent of the vector (R_1, \dots, R_p)

and is distributed as Z_{p+1} .

(iii) the law of V_{p+1} conditionally on R_1, \dots, R_p , and X_{p+1} ,
is characterized by (10).

Remarks:

a) The statements of the theorem should be presented somewhat differently since, e.g.: the property (i) by itself suffices to characterize the law of the sequence $(V_n, n \geq 1)$.

b) We could (should?) modify property (ii) as:

(ii') the increments: $X_1, X_2 - X_1, \dots, X_{p+1} - X_p, \dots$ are independent and distributed exponentially, with parameter 1.

c) The identity (10) admits many variants, some of which are the following: to relate efficiently formula (10) with Michael's formulae, it is interesting to give:

$$E \left[\exp \left(-\frac{x^2}{2V_1} \right) \middle| R_p, X_1 = \xi \right],$$

which I hope to discuss in the next document....