

↑ Juillet.

Quelques formules de grossissement initial ($1^{\text{er}} \rightarrow \text{expos'}$).

la formule générale (Thm 12.1, p. 33-34, Z II):

Si $\hat{\lambda}_t(\mathrm{d}l) = \lambda_t(\mathrm{d}l) \rho(l, t)$, alors si (X_t) est une (\mathcal{F}_t) martingale, on a: $X_t = \tilde{X}_t + \int_0^t \rho(L, s) d\langle X, B \rangle_s$.

Exemple 1: $L = \int_0^\infty \varphi(s) dB_s$, $\int_0^\infty ds \varphi^2(s) < \infty$.

Alors, λ_t et $\hat{\lambda}_t$ sont faciles à calculer, et on trouve:

$$\rho(l, s) = \varphi(s) \frac{l - m_s}{\sigma_s^2}, \text{ où: } m_s = \int_0^s \varphi(u) dB_u; \quad \sigma_s^2 = \int_s^\infty \varphi^2(u) du.$$

Cas particulier: $\varphi(s) = 1_{[0, t_0]}(s)$; i.e. $L = B_{t_0}$.

On obtient ainsi: $\rho(l, s) = \frac{l - B_s}{(t_0 - s)}, \quad s < t_0$.

et on a donc: $\tilde{B}_t = \tilde{B}_t + \int_{t_0}^t ds \frac{B_{t_0} - B_s}{t_0 - s}$

Comme maintenant (\tilde{B}_t) est indépendant de $\sigma(B_{t_0})$, on peut conditionner
 ∵ $B_{t_0} = y$ (par ex: 0), et on trouve ainsi la décomp. du front:

$$\tilde{B}_t = a + \tilde{B}_t + \int_0^t ds \frac{y - B_s}{(t_0 - s)}$$

À la suite, on se servira bcp. de la conséquence suivante pour le point de Bessel : $\mathbb{T}_{x \rightarrow 0}^{(t_0)}$:

$x \rightarrow 0$
ou r plutôt :

$$r(t) = r + \beta t + \frac{m-1}{2} \int_0^t \frac{ds}{r(s)} - \int_0^t \frac{ds r(s)}{(t-s)}.$$

En fait, cela nous mène au second Exemple :

Exemple 2: L'évolution de la filtration d'un processus de Markov avec X_t .

Sur un espace canonique, on a avec F_u (\tilde{F}_u) mesurable, on a :

$$(1) E_x[F_u | X_1 = y] = E_x[F_u \frac{p_{1-u}(X_u, y)}{p_1(x, y)}]$$

et donc, une (\tilde{F}_u) martingale (M_u) devient, pour la loi du point : $\mathbb{T}_{x \rightarrow y}^1$,

$$M_u = \tilde{M}_u + \int_0^u \frac{d \langle p_{1-u}(X_u, y), M \rangle_u}{p_{1-u}(X_u, y)}$$

Par ex, pour un mrt. b. avec drift : $\boxed{X_t = B_t + \int_0^t b(X_u) du}$, on obtient :

$$B_t = \tilde{B}_t + \int_0^t \frac{\partial}{\partial x} (\log p_{1-u}(X_u, y)) du,$$

et donc, pour le point : $\mathbb{T}_{x \rightarrow y}^1$, le drift b est changé en :

$$b_y(x) = b(x) + \frac{\partial}{\partial x} (\log p_{1-u}(x, y)).$$

Démonstration de la formule (1) :

$$\begin{aligned} E_x[f(X_1) F_u] &= E_x[f(X_1) E_x[F_u | X_1]] \\ &= E_x[P_{1-u} f(X_u) F_u] \\ &= E_x[F_u \int dy p_{1-u}(X_u, y) f(y)] \end{aligned}$$

$$\boxed{= \int dy p_1(x, y) f(y) E_x[F_u | X_1]}$$

Exemple 3: Décomposition du mouvement brownien
 $\{B_u, u \leq d_1\}$ sur $[0, g_1], [g_1, d_1], [g_1, 1]$, ...

- On commence par introduire l'op. de Brownian scaling;
- Puis, la décomposition: (faire un diagramme).

$$b_u = \frac{1}{\sqrt{g_t}} B_{g_t u}, u \leq 1, \quad m_u = \frac{1}{\sqrt{t-g_t}} |B_{g_t+u(t-g_t)}|, u \leq 1;$$

$$e_u = r_u = \frac{1}{\sqrt{dt-g_t}} |B_{g_t+u(t-g_t)}|, u \leq 1.$$

Cet exemple 3 sera décomposé en 2 parties :

(3.1) Mouvement initial avec $L = g = g_1$.

Les formules pour $\lambda_s(dl)$ et $\dot{\lambda}_s(dl)$ sont : $e(s, l)$

$$\lambda_s(dl) = \varepsilon_{g_s}(dl) \oplus \left(\frac{|B_s|}{\sqrt{1-s}} \right) + \frac{dl}{\pi} \underbrace{\begin{cases} 1 & s < l < 1 \\ \sqrt{(l-s)(1-l)} & \end{cases}}_{\text{III}} \exp \left(-\frac{B_s^2}{2(l-s)} \right)$$

d'où l'on déduit (mais cela utilise la formule du balayage ; à démontrer)

$$\dot{\lambda}_s(dl) = \varepsilon_{g_s}(dl) \oplus \left(\frac{|B_s|}{\sqrt{1-s}} \right) \frac{\operatorname{sgn}(B_s)}{\sqrt{1-s}} + dl e(s, l) \left(-\frac{B_s}{l-s} \right)$$

En conséquence, on a donc :

$$\rho(l, s) = \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|B_s|}{\sqrt{1-s}} \right) \frac{\operatorname{sgn}(B_s)}{\sqrt{1-s}} 1_{(\frac{s}{l} < 1)} + 1_{(l > s)} \left(-\frac{B_s}{l-s} \right) \quad (5)$$

La formule de l'érosionnement devient donc :

$$\beta_t = \tilde{\beta}_t + \int_0^t ds \left\{ 1_{(s < q)} \left(-\frac{B_s}{(q-s)} + 1_{(q \leq s)} \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|B_s|}{\sqrt{1-s}} \right) \right) \right\}$$

En d'autre forme, on trouve bien le pont brownien $(B_u, u \leq q)$, et le meandre pour $t \geq q$.

À ce point, il faut présenter la relation d'Imhof pour le meandre,

$$(3) \quad M = \frac{c}{X_1} \cdot P_0^{(3)}$$

En effet, ce qui se passe est que : $E_0^{(3)} \left(\frac{1}{X_1} \mid \mathcal{F}_t \right) = \psi \left(\frac{X_t}{\sqrt{1-t}} \right)$,
où : $\psi(x) = \frac{\Phi'(x)}{x}$

Alors, lorsque l'on écrit la formule de Chisanov pour le changt. en (3), on trouve bien :

$$(4) \quad X_t = \beta_t + \int_0^t \frac{ds}{X_s} + \int_0^t \frac{\psi'}{\psi} \left(\frac{X_s}{\sqrt{1-s}} \right) \frac{ds}{\sqrt{1-s}}$$

Or, $\frac{\psi'}{\psi}(x) = \frac{\Phi'}{\Phi} - \frac{1}{x}$, et donc de (4), on déduit bien que la formule de l'érosionnement ci-dessus est en accord....

(3.2) érosionnement initial avec $L = d \equiv d_1$.

Dans ce cas, on trouve :

$$(5) \lambda_\lambda(d) = \left(\frac{|z|}{\sqrt{2\pi(t-s)^3}} \exp \left(-\frac{x^2}{2(t-s)} \right) \right) \Big|_{x=B_s} 1_{(t>s)} dt.$$

On remarque bien sûr que ceci est :

$$\frac{y \exp(-y^2/2v)}{\sqrt{2\pi v^3}} = -\frac{\partial}{\partial y} (\mu_v(y)).$$

On va donc obtenir : $\rho(t, s) = \theta(v, x) \Big|_{\begin{array}{l} x=B_1 \\ v=t-s \end{array}} \quad 1(t \geq s),$

$$\text{avec : } \theta(v, y) = \frac{\frac{\partial^2}{\partial y^2} \mu_v(y)}{\frac{\partial}{\partial y} \mu_v(y)} = \frac{1}{x} - \frac{x}{v} \quad (x=y !)$$

Finalement, la formule de grossissement donne :

$$(6) \quad B_t = B_1 + \int_1^{t \wedge d} du \left\{ \frac{1}{B_u} - \frac{B_u}{(d-u)} \right\},$$

ce qui montre que le processus $\{B_{1+v}; v \leq d-1\}$

est un pont de Bessel 3, de longueur $(d-1)$.

Démonstration des formules (2) et (5) :

1)

Exemple 4: agrandissement de la filtration brownienne avec:

$$L = A_{\infty}^{(\nu)} = \int_0^{\infty} ds \exp 2(B_s + \sqrt{s}) , \quad \nu < 0; [\mu]$$

Pour le calcul de $\lambda_s^{(1)}(dl)$, il suffit seulement à l'aide de la connaissance de la loi de $A_{\infty}^{(\nu)}$, lorsque $B_0=0$, i.e.:

$$P(A_{\infty}^{(\nu)} \in dy) = P\left(A + \frac{1}{2} Z_{\mu} \in dy\right).$$

d'où:

$$\begin{aligned} \int \lambda_s^{(1)}(dl) f(l) &= E[f(A_{\infty}^{(\nu)}) | \mathcal{F}_1] = E[f(A_s^{(\nu)} + A_{\infty}^{(\nu)} \circ \theta_s) | \mathcal{F}_1] \\ &= E_{B_s^{(\nu)}} [f(A_s^{(\nu)}(\omega) + \tilde{A}_{\infty}^{(\nu)})]. \end{aligned}$$

On trouve alors, assez facilement, en posant:

$$A_t^{(\nu)} = \int_0^t ds \exp 2(B_s + \sqrt{s}) ; \quad r_t = \exp(B_t + \sqrt{t}) :$$

$$f(s, t) = r_t \frac{\partial}{\partial r} (\log \Phi)(r_t, A_t^{(\nu)}; t), \quad [\lambda \equiv]$$

ou :

$$\Phi(r, a; t) = \left(\frac{r^2}{2}\right)^{\mu} \frac{1}{(t-a)^{\mu+1}} \exp\left(-\frac{r^2}{2(t-a)}\right)$$

En conséquence,

$$\rho(s, t) = (2\mu) - \frac{r_s^2}{t - A_s^{(1)}}$$

- la formule de grossissement pour le mvt. brownien affirme :

$$B_t = \beta_t + \int_0^t ds \left\{ (2\mu) - \frac{\exp 2(B_s + \gamma_s)}{(A_\infty^{(1)} - A_s^{(1)})} \right\}$$

$$= \beta_t + \int_0^t ds \left\{ -(2\gamma) - \left(\exp 2(B_s + \gamma_s) \right) \left/ \int_1^\infty du \exp 2(B_u + \gamma_u) \right. \right\}$$

Vérifications: a) On peut s'assurer que l'espérance conditionnelle donne bien 0.

$$\text{i.e.: } -(2\gamma) - E\left(\frac{1}{A_\infty^{(1)}}\right) = -2\gamma - E[2\gamma Z_{(-)}] = -2\gamma + 2\gamma = 0.$$

b) Transformons le résultat précédent en un résultat sur les processus de Bond, au moyen de la rcp. de Lamperti, i.e.:

$$\exp(B_t + \gamma_t) = R^{(1)} \int_0^t du \exp 2(B_s + \gamma_s)$$

Le grossi^e précédent revient à grossir la filtration de $R^{(1)}$ avec $T_0^{(1)} = \inf\{t : R^{(1)}_t = 0\}$. Cela nous donne maintenant :

$$R_n^{(1)} = 1 + \gamma_u + \frac{2\mu+1}{2} \int_0^u \frac{du}{R_s^{(1)}} - \int_0^u ds \frac{R_s^{(1)}}{(T_0 - s)}$$

Autrement dit, $(R_u^{(1)}, u \leq T_0)$ est un pont de Bessel d'indice μ issu de 1, allant en 0, sur l'intervalle T_0 .

Exemple 5: Yosizawa de la ~~filt.~~ filt. brownienne avec tt la tribu du temps local.

Thm 12.2: $B_t = B_t^{\text{loc}} + \int_0^t ds \left\{ \frac{1}{B_s} - \frac{B_s}{(\delta_s - s)} \right\}. \quad (\text{p. 38-39-40})$

$$\left. \begin{array}{l} \frac{d}{dt} \\ t=0 \end{array} \right\}.$$

8th of July.

Second Lecture: Progressive enlargements / Martingales vanishing on the zero set

II. Martingales vanishing on the zero set. $Z = \{(t, \omega); t \leq 1, B_t(\omega) = 0\}$

Main thm: To $X \in L^1(\mathcal{F}_1)$, we associate: $X_t = E(X|\mathcal{F}_t)$.

Then, the following are equivalent:

$$1) X \in \mathcal{M}_1^0 ; 2) E(X|\mathcal{F}_g) = 0 ; 3) X_g = 0.$$

II.1. Preparation's.

a) the balayage formula: It will give us a nb. of examples of elements of \mathcal{M}_1^0 , but its applicability is much wider -

Proposition: If (X_t) is a semimartingale which vanishes on Z , then:

$$\int_0^t X_s ds = \int_0^t \int_{\mathcal{G}_s} dX_s.$$

In particular if $(X_t) \in \mathcal{M}_1^0$, then: $(\int_0^t X_s ds)$ is also in \mathcal{M}_1^0 .

Other examples: $f(l_t)|B_t| = \int_0^t f(l_s) d|B_s|$

Hence, $f(l_t)|B_t| - F(l_t)$ is a martingale.

Lemma 1: $E[\int_0^1 g_u] = E\left[\int_0^1 dL_u g_u\right]$, where: $dL_u = \frac{du}{\sqrt{1-u}} \sqrt{\frac{2}{\pi}}$

Proof: We have already seen this result, see above, and also in the 1st lecture. But, it can also be understood as a consequence of the balayage formula:

$$E[\int_0^1 g_u |B_1|] = E\left[\int_0^1 dL_u g_u\right],$$

and: $E[|B_1| |\mathcal{F}_g] = \sqrt{1-g_1} E[m_1]$ / etc. /.

zero set $\{\omega = 0\}$ of BM.

We shall see more such results later... (3rd lecture, e.g.).

Lemma 2: $E[\gamma_g X_g] = E[X \int_0^1 dL_u \gamma_u] \Rightarrow$

$$= E[E(X|\mathcal{F}_g) \int_0^1 dL_u]$$

Proof: The 1st equality follows from lemma 1, and integration by parts.

Question: Que se passe-t-il si (X_t) n'est pas un semimartingale? Puisque

$$E[\gamma_{gt} X_{gt}] = \text{etc...} ?$$

, then: II.2. Proof of the theorem. (p. 67).

II.3. Some examples of elements of \mathcal{M}^0 ; a general representation theorem.

We now know that elements of \mathcal{M}^0 are martingales whose terminal value ($\in \mathbb{R}$) say) is in the orthogonal of $L^2(\mathcal{F}_g)$, i.e.:

$$Y = X - E[X|\mathcal{F}_g].$$

This gives a systematic way to construct elements of \mathcal{M}_1^0 .

example:

$$\begin{aligned} X &= f(B_1), \quad E[X|\mathcal{F}_g] = E[f(\sqrt{1-g} \cdot m_1; \varepsilon)|\mathcal{F}_g] \\ &\stackrel{?}{=} \frac{1}{2} \int_0^\infty dy \frac{y}{(1-g)} \exp\left(-\frac{y^2}{2(1-g)}\right) (f(y)) \end{aligned}$$

Il paraît, on fait le calcul de:

$$Y_t = P_{1-t} f(B_t) - \frac{1}{2} \int_0^\infty dy y (f(y) + f'(y)) E[$$

or, on a déjà calculé à gauche de quantité' ...

Puis, la discussion amène aux questions d'équation conditionnelle ...

III. Some remarks about \mathcal{F}_{L^-} and \mathcal{F}_{L^+} .

By definition: $\mathcal{F}_{L^-} = \sigma\{\mathcal{Z}_L; \mathcal{Z}(\mathcal{F}_t) \text{ predictable}\}$

$\mathcal{F}_{L^+} = \sigma\{\mathcal{Z}_L; \mathcal{Z}(\mathcal{F}_t) \text{ prog. measurable}\}$

$\mathcal{Z}(\mathcal{F}_t) \text{ prog. measurable sets / process ... are complicated, but ... }]$

Now, since L is a (\mathcal{F}_t) stopping time, we had also the more classical:
etc... $\mathcal{Z}(\mathcal{F}_t^L)$, and:

$$\mathcal{F}_{L^+} = (\mathcal{F}^{\text{prog}, L})_L; \quad \mathcal{F}_{L^+-} = \mathcal{F}_{L-}^{\text{prog}, L}.$$

How can we compute cond. expectations w.r.t. \mathcal{F}_{L^-} and/or \mathcal{F}_{L^+} ?

$$E[X | \mathcal{F}_{L^-}] = \lim_{u \uparrow L} \frac{E[X 1_{(u < L)} | \mathcal{F}_u]}{P(L > u | \mathcal{F}_u)}$$

$$E[X | \mathcal{F}_{L^+}] = \lim_{v \downarrow 0} \frac{E[X 1_{(L < v)} | \mathcal{F}_v]}{(1 - P(L > v))} \quad v = L + u$$

Is \mathcal{F}_{L^+} diff. from \mathcal{F}_{L^-} ?

Theorem: Let $L = \sup\{t: M_t = 0\}$, where (M_t) is unif. integrable, and $P(M_\infty = 0) = 0$. Then,
 \mathcal{F}_{L^+} strictly contains \mathcal{F}_L ; in fact, $\text{sgn}(M_\infty)$ is \mathcal{F}_{L^+} measurable,
and not in \mathcal{F}_L . Proof: $E[M_\infty | \mathcal{F}_{L^-}] = 0$; but, $E[|M_\infty| | \mathcal{F}_{L^-}] \neq$

In the particular case of Brownian motion, we can show: $\mathcal{F}_{B^+} = \mathcal{F}_B \vee \sigma(\text{sgn}(B))$

I. Formule de grossissement progressif.

Defn. d'un temps honeste: $L = \inf_t L_t$, sur $(L \leq t)$.

Fins d'ensembles prévisibles: $L = \sup \{ u : (u, \omega) \in \Gamma \}$.

Si L est honnête, alors on définit $\mathcal{F}_t^{\text{gross}}_L = \mathcal{F}_t^L$ pour l'instant / Bien fini, mais je préciserai....

Lemma 1: Si (H_u) est un proc. prévisible / (\mathcal{F}_t^L) , alors: $H_u = H'_u 1_{(u \leq L)} + H''_u 1_{(u > L)}$ avec deux processus (\mathcal{F}_t) prévisibles H' et H'' .

→ Ceci va nous permettre d'obtenir la formule de décomposition, i.e.:
à L , on ait: $Z_t^L = P(L > t | \mathcal{F}_t) = M_t^L - A_t^L$.

Théorème: $X_t = \tilde{X}_t + \int_0^{t \wedge L} \frac{d\langle X, Z_u^L \rangle}{Z_u^L} + \int_L^t \frac{d\langle X, 1 - Z_u^L \rangle}{(1 - Z_u^L)}$

To prove the theorem, we first need an easy: / rappel: $Z_t^L = M_t^L - A_t^L$

Lemma 2: a) For every $\gamma \geq 0$, prévisible process (γ_u) , one has:

$$E[\gamma_L] = E\left[\int_0^\infty dA_u^L \gamma_u\right]$$

b) For every uniformly integrable martingale (X_t) , one has

$$E[X_L] = E[X_\infty A_\infty^L] = E[X_\infty M_\infty^L].$$

$$\left[\mathcal{F}_L \right] \neq 0.$$

Bartlow's conjecture: $\mathcal{F}_{L+} = \mathcal{F}_L \vee \sigma(A)$.

Proof of Lemma 2: It suffices to prove the 1st formula for elementary g^L_s : $g_u = 1_{[T]} 1_{[t, \infty]} (\frac{z}{z_u})$.

then:

$$\begin{aligned} E[g_L] &= E[1_{[T]} z_T^L] = -E[1_{[T]} (z_\infty^L - z_T^L)] \\ &= E[1_{[T]} (A_\infty^L - A_T^L)] \\ &= E\left[\int_0^\infty dA_u^L g_u\right]. \end{aligned}$$

For the 2nd formula, we use the first result, and integration by parts.

Proof of the theorem: We want to show:

$$E\left[\int_0^\infty H_u dX_u\right] = E\left[\int_0^\infty H_u \left\{ \underbrace{\frac{d\langle X, Z \rangle_u^L}{Z_u^L}}_{1(L < u)} + \underbrace{\frac{d\langle X, 1-Z \rangle_u^L}{(1-Z_u^L)}}_{1(L > u)} \right\}\right]$$

when (H_u) is a simple (F_u^L) pred. process, i.e. H^L & H^U are simple.
We start from the RHS

$$\begin{aligned} &= E\left[\int_0^L H_u^L \frac{d\langle X, Z \rangle_u^L}{Z_u^L}\right] + E\left[\int_L^\infty H_u^U \left(\frac{d\langle X, 1-Z \rangle_u^L}{1-Z_u^L}\right)\right] \\ &= E\left[\int_0^\infty H_u^L d\langle X, Z \rangle_u^L\right] + E\left[\int_0^\infty H_u^U d\langle X, 1-Z \rangle_u^L\right] \\ &= E\left[\int_0^\infty H_u^L d\langle X, M \rangle_u^L\right] - E\left[\int_0^\infty H_u^U d\langle X, M \rangle_u^L\right] \end{aligned}$$

$$= E \left[\int_0^L H'_u dX_u - \int_0^L H''_u dX_u \right]$$

$$= E \left[\int_0^L H'_u dX_u + \int_L^\infty H''_u dX_u \right] = E \left[\int_0^\infty H_u dX_u \right].$$

We then consider our previous example with $L = g_1$, but now we have 2 enlarged formulae:

1st enf. formula [in $(\mathcal{F}_t^{prog, g_1})$]:

$$\text{Recall, } Z_u^{g_1} = P(g_1 > u | \mathcal{F}_u) \\ = \frac{1}{\Phi} \left(\frac{|Bu|}{\sqrt{1-u}} \right)$$

$$\beta_t = \beta_t^{(prog)} + \int_0^{t \wedge g_1} \left(-\frac{\Phi'}{\Phi} \right) \left(\frac{|Bu|}{\sqrt{1-u}} \right) \frac{\text{sgn}(Bu)}{\sqrt{1-u}} du$$

$$\mathcal{F}_t^{prog, g_1} \subset \mathcal{F}_t^{in, g_1} / + \int_{g_1}^t \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|Bu|}{\sqrt{1-u}} \right) \frac{\text{sgn}(Bu)}{\sqrt{1-u}}$$

2nd enf. formula: [in $(\mathcal{F}_t^{in, g_1})$]:

$$\beta_t = \beta_t^{(in)} - \int_0^{t \wedge g_1} \frac{du \ Bu}{(g_1 - u)} + \int_{g_1}^t \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|Bu|}{\sqrt{1-u}} \right) \frac{\text{sgn}(Bu)}{\sqrt{1-u}} du$$

Thus, we obtain:

$$\beta_t^{(in)} = \beta_t^{(prog)} + \int_0^{t \wedge g_1} \frac{du \ Bu}{(g_1 - u)} - \int_0^{t \wedge g_1} \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|Bu|}{\sqrt{1-u}} \right) \frac{\text{sgn}(Bu)}{\sqrt{1-u}} du$$

Pour que ces calculs soient compatibles, il faut donc vérifier montrer:

$$E \left[\frac{1(u < g_1)}{(g_1 - u)} \mid \mathcal{F}_u^{prog, g_1} \right] |Bu| = \left(\frac{\Phi'}{\Phi} \right) \left(\frac{|Bu|}{\sqrt{1-u}} \right) \frac{1}{\sqrt{1-u}}$$

qui découle d'une formule explicite (lemme 3, dans le 1^{er} exposé).

Details.].

Un exemple, avec dernier temps de passage:

$(t, t \geq 0)$ (\mathcal{F}_t) local mart^{*}; ≥ 0 ; $Y_t \xrightarrow[t \rightarrow \infty]{} 0$;

$< Y_0$; $\gamma_y = \sup\{u \geq 0 : Y_u = y\}$.

we have the following

Lemma: i) $\sup_{t \geq 0} Y_t \stackrel{\text{(aw)}}{=} Y_0$

ii) $Z_t^y = P(\gamma_y > u | \mathcal{F}_u) = \left(\frac{Y_u}{y}\right) \wedge 1.$

$$= Z_0^y + \frac{1}{y} \int_0^u 1(Y_v \leq y) dY_v - \frac{1}{2y} \ell_t^y.$$

the enlarg. formula becomes:

$$X_t = \tilde{X}_t + \int_0^{t \wedge \gamma_y} \frac{d\langle X, Y \rangle_v 1(Y_v \leq y)}{Y_v} - \int_{\gamma_y}^t \frac{d\langle X, Y \rangle_v}{(y - Y_v)}.$$

In particular, $\underbrace{Y_{\gamma_1+v} - y}_{\leq 0} = (\tilde{Y}_{\gamma_1+v} - \tilde{Y}_y) - \int_0^{t+v} \frac{d\langle Y_{v+t}, y \rangle}{(y - Y_{v+t}y)}$

donc : $3 \dim ?? \text{ BES} /$
soudain bizarre ... / (Vérifier)

En ce cas, lorsque l'on applique ceci à $Y_t = 1/R_t$, avec $R = \text{BES}(3)$
il vient :

$$\underbrace{\{R, Y_1 + t - 1; t \geq 0\}}_{: \text{BES}(3)} : \text{BES}(3).$$

9th of July.A third note for Prof. Fujita.

At the end of the Pitman-Yor paper: "On the lengths of excursions..." (Sem. Prob. XXXI,)

We have a discussion of generalized arcsine laws, which may applied again to yield some description of the joint law of (X_1, A_1^+) ,

where (X_t) is a skew Bessel process, with dimension $2-\alpha$, and skewness parameter μ ; we denote the law of this process by $P_{\alpha, \mu}$. Then, we can write the following:

$$\begin{aligned} E_{\alpha, \mu} \left[f(X_1) \frac{1}{(1+\lambda A_1^+)^{\alpha}} 1(X_1 > 0) \right] &\stackrel{\text{def}}{=} \pi_{\alpha, \mu}(f, \lambda). \\ &= E_{\alpha, \mu} \left[f(\sqrt{1-q} m_1) \frac{1}{(1+\lambda [q A_{br}^+ + (1-q)])^{\alpha}} \right] p \end{aligned}$$

where A_{br}^+ denotes the time spent ≥ 0 by the bridge.

Then, we also know the following:

i) q is beta($\alpha, 1-\alpha$) distributed, this result goes back to Dynkin (1961)

ii) the law of m_1 does not depend on α (this is a remarkable feature), i.e.

$$P_{\alpha}(m_1 \in dp) = p e^{-p^2/2} dp.$$

iii) for every $\mu \geq 0$,

$$E_{\alpha, \mu} \left[\frac{1}{(1+\mu A_{br}^+)^{\alpha}} \right] = \frac{1}{\mu(1+\mu)^{\alpha} + q}.$$

$$\text{If we write: } \Pi_{\alpha, p}(f, \lambda) = p E_{\alpha, p} \left[f(\sqrt{1-q} m_1) \frac{1}{D^\alpha} \right],$$

where:

$$D = 1 + \lambda(1-q) + \lambda q A_{br}^+ = (1 + \lambda(1-q)) \left(1 + \mu(q) A_{br}^+ \right),$$

where: $\mu(q) = \frac{\lambda q}{1 + \lambda(1-q)}$, we now deduce from iii), that:

$$\Pi_{\alpha, p}(f, \lambda) = E_{\alpha, p} \left[f(\sqrt{1-q} m_1) \frac{p}{(1 + \lambda(1-q))^\alpha [p(1 + \mu(q))^\alpha + q]} \right]$$

Since:

$$1 + \mu(q) = \frac{1 + \lambda}{1 + \lambda(1-q)}, \text{ we finally obtain:}$$

$$\Pi_{\alpha, p}(f, \lambda) = E_{\alpha, p} \left[f(\sqrt{1-q} m_1) \frac{p}{p(1 + \lambda)^\alpha + q(1 + \lambda(1-q))^\alpha} \right]$$

Now, it seems plausible that, for $a > 0$, there exists a r.v. $X_{p,a}$ such that:

$$E \left[\frac{1}{(1 + \lambda X_{p,a})^\alpha} \right] = \frac{1}{p(1 + \lambda)^\alpha + q(1 + \lambda a)^\alpha}$$

Assuming this is the case, we have obtained:

$$\Pi_{\alpha, p}(f, \lambda) = E_{\alpha, p} \left[f(\sqrt{1-q} m_1) \frac{p}{(1 + \lambda X_{p,1-q})^\alpha} \right],$$

and finally, we find that, conditionally on ($X_1 > 0$),

$$(X_1, A_1^+) \stackrel{(law)}{=} (\sqrt{1-q} m_1, X_{p,1-q}).$$