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Some formulae for the Brownian and Bessel spiders, II

1)

The technique used to obtain the formulae in I is essentially excursion theory [in fact, starting from independent exponential time, and using scaling, to arrive to functionals taken at $T_* = T(x_1, \dots, x_k)$]

Here, I will prove all the formulae in I, and more, using stochastic calculus for the spider - I look only at the Brownian case, the Bessel case being essentially similar.

1. Warm up: A proof of (9)

$$P(B_{T_*} = x_i) = \frac{(1/a_i)}{\sum_l (1/a_l)}$$

Define: $B_t^{(i)} = |B_t| 1(B_t \in I_i)$; then, $(B_t^{(i)} - \frac{1}{k} L_t; t \geq 0)$

is a martingale. Hence, we have: (9.a) $E[B_{T_*}^{(i)}] = \frac{1}{k} E[L_{T_*}]$.

Call C the right-hand side, which does not depend on (i) ; so, we can write (9.a) in the form:

$$(9.a') \quad a_i P(B_{T_*} = x_i) = C,$$

and, therefore: (9.b) $P(B_{T_*} = x_i) = \left(\frac{C}{a_i}\right)$.

To compute C , it remains to sum up the different quantities on the left-hand side of (9.b), and we obtain:

$$1 = C \left(\sum_i \frac{1}{a_i} \right).$$

Hence, $C = 1 / \sum_i (1/a_i)$, and, going back to (9.b), we obtain (9).

2. A proof of (7), (7)_i and (10).

I use the same idea, but now with the semimartingale:

$$M_t^{(i)} \stackrel{\text{def}}{=} \sinh(\lambda B_t^{(i)}) \exp\left(-\frac{\lambda^2 t}{2}\right)$$

Itô's formula tells us that:

$$M_t^{(i)} = \lambda \int_0^t \cosh(\lambda B_s^{(i)}) e^{-\frac{\lambda^2 s}{2}} dB_s^{(i)} = \text{mart} + \left(\frac{\lambda}{k}\right) \int_0^t e^{-\frac{\lambda^2 s}{2}} dL_s$$

Hence, we obtain again:

$$(7.a) \quad E \left[\sinh(\lambda B_{T_*}^{(i)}) \exp\left(-\frac{\lambda^2 T_*}{2}\right) \right] = \frac{\lambda}{k} E \left[\int_0^{T_*} \exp\left(-\frac{\lambda^2 s}{2}\right) dL_s \right]$$

so that we find that the left-hand side does not depend on (i); call this common value C again; we have:

$$(7.a') \quad \sinh(\lambda a_i) E \left[\exp\left(-\frac{\lambda^2 T_*}{2}\right) 1_{(B_{T_*} = a_i)} \right] = C$$

Now, in order to compute C, we remark that:

$$(7.b) \quad E \left[\cosh(\lambda B_{T_*}) \exp\left(-\frac{\lambda^2 T_*}{2}\right) \right] = 1.$$

Developing the left-hand side of (7.b), we get:

$$(7.c) \quad \left(\sum_i \coth(\lambda a_i) \right) C = 1,$$

so that, now, we can plug this result in (7.a), and we get:

$$(7)_i \quad E \left[\exp\left(-\frac{\lambda^2 T_*}{2}\right) 1_{(B_{T_*} = a_i)} \right] = \frac{(1 / \sinh(\lambda a_i))}{\sum_l \coth(\lambda a_l)}$$

3. Further amplification and proof of (10) and (11).

Let $(z_u; u \geq 0)$ be a bounded predictable process, and $q_t = \sup\{s \leq t; B_s = 0\}$

We have:

$$\begin{aligned} z_{q_t} M_t^{(i)} &= z_{q_t} \left(\sinh(\lambda B_t^{(i)}) \exp\left(-\frac{\lambda^2 t}{2}\right) \right) \\ &= \int_0^t z_{q_s} dM_s^{(i)}, \quad \text{from the Itô formula.} \\ &= (\text{mart}) + \frac{\lambda}{k} \int_0^t z_{q_s} \exp\left(-\frac{\lambda^2 s}{2}\right) dL_s. \end{aligned}$$

Hence, we have:

$$(12.a) \quad E \left[z_{q_{T_*}} \sinh(\lambda B_{T_*}^{(i)}) e^{-\frac{\lambda^2}{2} T_*} \right] = \frac{\lambda}{k} E \left[\int_0^{T_*} e^{-\frac{\lambda^2}{2} s} z_{q_s} dL_s \right],$$

and again this common value does not depend on (i) .

Let me change z into $\tilde{z}_s = z_{q_s} \exp\left(-\frac{\lambda^2}{2} s\right)$, and write: $\tilde{T}_* = T_* - q_{T_*}$.

Now, we can write (12.a) as:

$$(12.a') \quad E \left[\tilde{z}_{q_{T_*}} \sinh(\lambda a_i) e^{-\frac{\lambda^2}{2} \tilde{T}_*} 1_{(B_{T_*} = a_i)} \right] = \frac{\lambda}{k} E \left[\int_0^{\tilde{T}_*} \tilde{z}_{q_s} dL_s \right]$$

so that:

$$(12.a'') \quad E \left[\tilde{z}_{q_{T_*}} e^{-\frac{\lambda^2}{2} \tilde{T}_*} 1_{(B_{T_*} = a_i)} \right] = \left(\frac{\lambda}{\sinh(\lambda a_i)} \right) \frac{1}{k} E \left[\int_0^{\tilde{T}_*} \tilde{z}_{q_s} dL_s \right]$$

From there, we obtain many results all at once; in fact, we see a good fraction of

the assertions a) and b), on p. 2) of I.

Reading from (12. a''), we get:

a) $\mathcal{F}_{g_{T_*}}$ and the pair $\{\tilde{T}_*, B_{T_*}\}$ are independent;

$$b) E \left[\exp\left(-\frac{\lambda^2}{2} \tilde{T}_*\right) 1_{(B_{T_*} = \alpha_i)} \right] = \frac{\lambda \alpha_i}{\sinh(\lambda \alpha_i)} \left(\frac{1}{k} E[L_{T_*}] \right)$$

$$= \left(\frac{\lambda \alpha_i}{\sinh(\lambda \alpha_i)} \right) \left(\frac{(1/\alpha_i)}{\sum_{\ell} (1/\alpha_{\ell})} \right),$$

from (9.a) and sequel.

Hence, we get both (9) and (10) this way:

$$(10) \quad E \left[\exp\left(-\frac{\lambda^2}{2} \tilde{T}_*\right) \mid B_{T_*} = \alpha_i \right] = \frac{\lambda \alpha_i}{\sinh(\lambda \alpha_i)}$$

$$c) \quad E \left[\tilde{z}_{g_{T_*}} \right] = \left\{ \frac{1}{k} E \left[\int_0^{T_*} \tilde{z}_{\Delta} dL_{\Delta} \right] \right\} \left\{ \frac{1}{k} E[L_{T_*}] \right\}$$

in other words, the dual previsible projection of $1_{(g_{T_*} \leq t)}$ is $\frac{1}{k} (L_{t \wedge T_*}) / \frac{1}{k} E[L_{T_*}]$.

[Note that the developments in this third part are sensibly different from those in the 2 first parts]

In particular, there is no exact computation of the constant C /.